

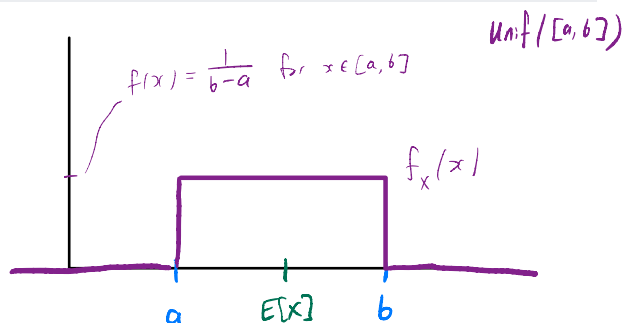
W06 Notes

Continuous families: summary

01 Theory

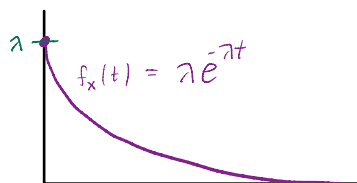
Uniform: $X \sim \text{Unif}([a, b])$

- All times $a \leq t \leq b$ equally likely.
- $f_X(t) = \frac{1}{b-a}$
- $E[X] = \frac{a+b}{2}$
- $\text{Var}[X] = \frac{1}{12}(b-a)^2$



Exponential: $X \sim \text{Exp}(\lambda)$

- Measures wait time until first arrival.
- $f_X(t) = \lambda e^{-\lambda t}$
- $E[X] = \frac{1}{\lambda}$
- $\text{Var}[X] = \frac{1}{\lambda^2}$



Erlang: $X \sim \text{Erlang}(\ell, \lambda)$

- Measures wait time until ℓ^{th} arrival.
- $f_X(t) = \frac{\lambda^\ell}{(\ell-1)!} t^{\ell-1} e^{-\lambda t}$
- $E[X] = \frac{\ell}{\lambda}$
- $\text{Var}[X] = \frac{\ell}{\lambda^2}$

We will see later:

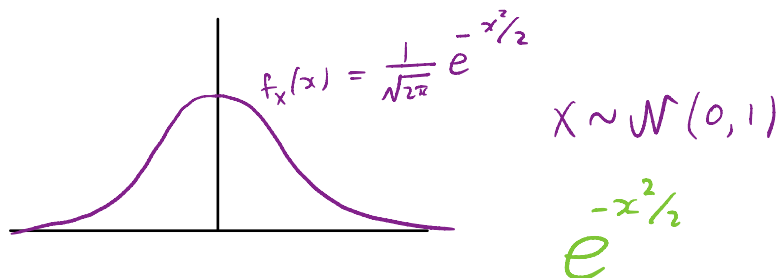
$$\text{Exp}(\lambda) + \text{Exp}(\lambda) = \text{Erlang}(2, \lambda)$$

$$\text{ie. } X, Y \sim \text{Exp}(\lambda)$$

$$\text{then } X+Y \sim \text{Erlang}(2, \lambda)$$

Normal: $X \sim \mathcal{N}(\mu, \sigma^2)$

- Limiting distribution of large sums.
- $f_X(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-(x-\mu)^2/2\sigma^2}$
- $E[X] = \mu$
- $\text{Var}[X] = \sigma^2$



Normal distribution

02 Theory

□ Normal distribution

A variable X has a **normal distribution**, written $X \sim \mathcal{N}(\mu, \sigma^2)$ or “ X is Gaussian (μ, σ) ,” when it has PDF given by:

$$f_X(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-(x-\mu)^2/2\sigma^2}$$

The **standard normal** is $Z \sim \mathcal{N}(0, 1)$ and its PDF is **usually denoted by** $\varphi(x)$:

$$\varphi(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}$$

The standard normal CDF is **usually denoted by** $\Phi(z)$:

$$\Phi(z) = \int_{-\infty}^z \frac{1}{\sqrt{2\pi}} e^{-u^2/2} du$$

“standard”
 $:= \begin{cases} \mu = 0 \\ \sigma = 1 \end{cases}$

- To show that $\varphi(x)$ is a valid probability density, we must show that $\int_{-\infty}^{+\infty} \varphi(x) dx = 1$.
 - This calculation is *not trivial*; it requires a double integral in polar coordinates!
- There is *no explicit antiderivative* of φ
 - A computer is needed for numerical calculations.
 - A *chart of approximate values* of Φ is provided for exams.

- To check that $E[Z] = 0$:
 - Observe that $x\varphi(x)$ is an *odd function*, i.e. symmetric about the y -axis.
 - One must then simply verify that the improper integral converges.
- To check that $\text{Var}[Z] = 1$:
 - Since $\mu = E[Z] = 0$, we find:

$$\text{Var}[Z] = E[Z^2] \gg \gg \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{+\infty} x^2 e^{-x^2/2} dx =: I$$

$\xrightarrow{\text{Handwritten: } \text{Var}[Z] = E[Z^2] - E[Z]^2 \rightarrow 0}$
 $\xrightarrow{\text{Handwritten: } \int x^2 e^{-x^2/2} dx \text{ with } u=x, dv=e^{-x^2/2} dx}$

- Use integration by parts to compute that $I = 1$. (Select $u = x$ and $dv = x e^{-x^2/2} dx$.)

General and standard normals

Assume that $Z \sim \mathcal{N}(0, 1)$ and σ, μ are constants. Define $X = \sigma Z + \mu$. Then:

$$f_X = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-(x-\mu)^2/2\sigma^2}$$

That is, $\sigma Z + \mu$ has the distribution type $\mathcal{N}(\mu, \sigma^2)$.

Derivation of PDF of $\sigma Z + \mu$ >

Suppose that $X = \sigma Z + \mu$. Then:

$$\begin{aligned}
 F_X(x) &= P[X \leq x] \\
 &= P[\sigma Z + \mu \leq x] \\
 &= P\left[Z \leq \frac{x-\mu}{\sigma}\right] = F_Z\left(\frac{x-\mu}{\sigma}\right) \\
 &= \Phi\left(\frac{x-\mu}{\sigma}\right)
 \end{aligned}$$

Differentiate to find f_X :

$$\begin{aligned}
 f_X(x) &= \frac{d}{dx} F_X(x) \\
 &= \frac{d}{dx} \Phi\left(\frac{x-\mu}{\sigma}\right) \quad \frac{d}{dx} \Phi = \varphi \\
 &= \frac{1}{\sigma} \varphi\left(\frac{x-\mu}{\sigma}\right) \quad e^{-x^2/2} \rightsquigarrow e^{-(x-\mu)^2/2} \\
 &= \frac{1}{\sqrt{2\pi}\sigma} e^{-(x-\mu)^2/2\sigma^2}
 \end{aligned}$$

From this fact we can infer that $E[X] = \mu$ and $\text{Var}[X] = \sigma^2$ whenever $X \sim \mathcal{N}(\mu, \sigma^2)$.

03 Illustration

Example - Basic generalized normal calculation

Suppose $X \sim \mathcal{N}(-3, 4)$. Find $P[X \geq -1.7]$.

Solution

First write X as a linear transformation of Z :

$$X \sim 2Z - 3$$

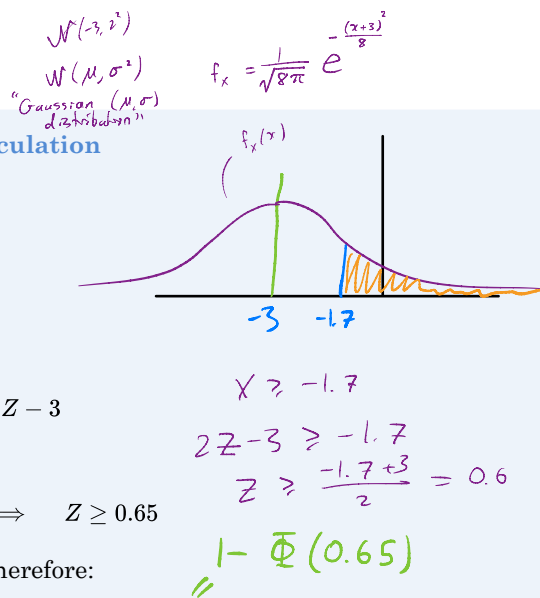
Then:

$$X \geq -1.7 \iff Z \geq 0.65$$

Look in a table to find that $\Phi(0.65) \approx 0.74$ and therefore:

$$P[Z \geq 0.65] \ggg 1 - P[Z \leq 0.65]$$

$$\ggg \approx 1 - 0.74 \ggg \boxed{0.26}$$



Example - Gaussian: interval of 2/3

Find the number a such that $P[-a \leq Z \leq +a] = 2/3$.

Solution

First convert the question:

$$P[-a \leq Z \leq +a] \ggg F_Z(a) - F_Z(-a)$$

$$\ggg \Phi(a) - \Phi(-a)$$

$$\ggg 2\Phi(a) - 1$$

Solve for a so that this value is $2/3$:

$$2\Phi(a) - 1 = 2/3 \ggg \Phi(a) = 5/6 \ggg a = \Phi^{-1}(5/6)$$

Use a Φ table to conclude $a \approx 0.97$.

≡ Example - Heights of American males

Suppose that the height of an American male in inches follows the normal distribution $\mathcal{N}(71, 6.25)$.

(a) What percent of American males are over 6 feet, 2 inches tall?

(b) What percent of those over 6 feet tall are also over 6 feet, 5 inches tall?

Solution

(a)

Let H be a random variable measuring the height of American males in inches, so $H \sim \mathcal{N}(71, 2.5^2)$. Thus $H \sim 2.5Z + 71$, and:

$$P[H > 74] \ggg 1 - P[H \leq 74]$$

$$\ggg 1 - P[2.5Z + 71 \leq 74]$$

$$\ggg 1 - P[Z \leq 1.20]$$

$$\ggg 1 - 0.8849 \approx 11.5\%$$

(b)

We seek $P[H > 77 \mid H > 72]$ as the answer. Compute as follows:

$$P[H > 77 \mid H > 72] = \frac{P[H > 77]}{P[H > 72]}$$

$$\ggg \frac{P[2.5Z + 71 > 77]}{P[2.5Z + 71 > 72]}$$

$$\ggg \frac{1 - P[Z \leq 2.4]}{1 - P[Z \leq 0.4]} = \frac{1 - 0.9918}{1 - 0.6554} \approx 2.38\%$$

≡ Example - Variance of normal from CDF table

Suppose $X \sim \mathcal{N}(5, \sigma^2)$, and suppose you know $P[X > 9] = 0.2$.

Find the approximate value of σ using a Φ table.

Solution

$$X \sim \mathcal{N}(5, \sigma^2) \implies X \sim \sigma Z + 5$$

So $1 - P[X \leq 9] = 0.2$ and thus $P[\sigma Z + 5 \leq 9] = 0.8$. Then:

$$P[\sigma Z + 5 \leq 9] = P[Z \leq 4/\sigma]$$

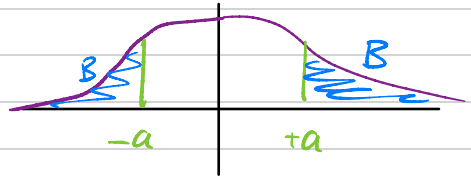
so $P[Z \leq 4/\sigma] = 0.8$.

Looking in the chart of Φ for the nearest inverse of 0.8, we obtain $4/\sigma = 0.842$, hence $\sigma = 4.75$.

Find the number a such that $P[-a \leq Z \leq +a] = 2/3$.

$$P[-a \leq Z \leq +a] = \int_{-a}^{+a} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx = 2/3$$

||
ans

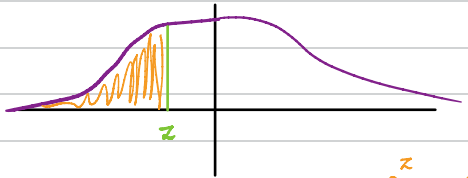


$$B = \Phi(-a)$$

$$\text{ans} = 1 - 2\Phi(-a)$$

$$B = 1 - \Phi(+a)$$

$$\begin{aligned} \text{ans} &= 1 - 2(1 - \Phi(a)) \\ &= 2\Phi(a) - 1 \end{aligned}$$



$$\text{chart gives } \Phi(z) = \int_{-\infty}^z \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx$$

$$\frac{2}{3} = 2\Phi(a) - 1$$

$$\Phi(a) = \frac{5}{6}$$

$$\leadsto \text{chart} \leadsto a = \boxed{0.97}$$

$$\text{because } \Phi(0.97) \approx 0.8340$$

$$\text{Lesson: } \underline{\Phi(-z) = 1 - \Phi(+z)}$$

Suppose that the height of an American male in inches follows the normal distribution $N(71, 6.25)$.

(a) What percent of American males are over 6 feet, 2 inches tall?

(b) What percent of those over 6 feet tall are also over 6 feet, 5 inches tall?

$$X \sim N(71, 2.5^2) \quad \mu = 71, \quad \sigma = 2.5$$

$$\begin{aligned} (a) \quad P[X > 74] &= P[2.5Z + 71 > 74] \\ &= P\left[Z > \frac{74 - 71}{2.5}\right] \quad \frac{\text{val} - \mu}{\sigma} \\ &= P[Z > 1.2] \quad = \text{"z-score"} \\ &= 1 - P[Z \leq 1.2] \approx 1 - 0.8849 \\ &= 11.5\% \end{aligned}$$

$$\begin{aligned} (b) \quad P[X > 77 \mid X > 72] &= \frac{P[X > 77 \cap X > 72]}{P[X > 72]} \\ &= \frac{P[X > 77]}{P[X > 72]} = \frac{P[2.5Z + 71 > 77]}{P[2.5Z + 71 > 72]} \\ &= \frac{1 - P[Z \leq 2.4]}{1 - P[Z \leq 0.9]} = \frac{1 - 0.9918}{1 - 0.6554} \approx 2.38\% \end{aligned}$$

Suppose $X \sim \mathcal{N}(5, \sigma^2)$, and suppose you know $P[X > 9] = 0.2$.

Find the approximate value of σ using a Φ table.

$$1 - 0.2 = P[X \leq 9] \\ 0.8$$

$$X \sim \sigma Z + 5$$

$$\ggg P[\sigma Z + 5 \leq 9] = 0.8$$

$$P[Z \leq 4/\sigma] = 0.8$$

$$\Phi(4/\sigma) = 0.8$$

$$4/\sigma = \Phi^{-1}(0.8) \\ = 0.842$$

$$\ggg \sigma = 4/0.842 = 4.75$$