

# W06 Notes

## Continuous families: summary

### 01 Theory

**Uniform:**  $X \sim \text{Unif}([a, b])$

- All times  $a \leq t \leq b$  equally likely.
- $f_X(t) = \frac{1}{b-a}$
- $E[X] = \frac{a+b}{2}$
- $\text{Var}[X] = \frac{1}{12}(b-a)^2$

**Exponential:**  $X \sim \text{Exp}(\lambda)$

- Measures wait time until first arrival.
- $f_X(t) = \lambda e^{-\lambda t}$
- $E[X] = \frac{1}{\lambda}$
- $\text{Var}[X] = \frac{1}{\lambda^2}$

**Erlang:**  $X \sim \text{Erlang}(\ell, \lambda)$

- Measures wait time until  $\ell^{\text{th}}$  arrival.
- $f_X(t) = \frac{\lambda^\ell}{(\ell-1)!} t^{\ell-1} e^{-\lambda t}$
- $E[X] = \frac{\ell}{\lambda}$
- $\text{Var}[X] = \frac{\ell}{\lambda^2}$

**Normal:**  $X \sim \mathcal{N}(\mu, \sigma^2)$

- Limiting distribution of large sums.
- $f_X(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-(x-\mu)^2/2\sigma^2}$
- $E[X] = \mu$
- $\text{Var}[X] = \sigma^2$

## Normal distribution

### 02 Theory

#### ◻ Normal distribution

A variable  $X$  has a **normal distribution**, written  $X \sim \mathcal{N}(\mu, \sigma^2)$  or “ $X$  is Gaussian  $(\mu, \sigma)$ ,” when it has PDF given by:

$$f_X(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-(x-\mu)^2/2\sigma^2}$$

The **standard normal** is  $Z \sim \mathcal{N}(0, 1)$  and its PDF is **usually denoted by  $\varphi(x)$** :

$$\varphi(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}$$

The standard normal CDF is **usually denoted by  $\Phi(z)$** :

$$\Phi(z) = \int_{-\infty}^z \frac{1}{\sqrt{2\pi}} e^{-u^2/2} du$$

- To show that  $\varphi(x)$  is a valid probability density, we must show that  $\int_{-\infty}^{+\infty} \varphi(x) dx = 1$ .
  - This calculation is *not trivial*; it requires a double integral in polar coordinates!
- There is *no explicit antiderivative* of  $\varphi$ 
  - A computer is needed for numerical calculations.
  - A *chart of approximate values* of  $\Phi$  is provided for exams.

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- To check that  $E[Z] = 0$ :
  - Observe that  $x\varphi(x)$  is an *odd function*, i.e. symmetric about the  $y$ -axis.
  - One must then simply verify that the improper integral converges.
- To check that  $\text{Var}[Z] = 1$ :
  - Since  $\mu = E[Z] = 0$ , we find:

$$\text{Var}[Z] = E[Z^2] \implies \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{+\infty} x^2 e^{-x^2/2} dx =: I$$

- Use integration by parts to compute that  $I = 1$ . (Select  $u = x$  and  $dv = xe^{-x^2/2} dx$ .)

### General and standard normals

Assume that  $Z \sim \mathcal{N}(0, 1)$  and  $\sigma, \mu$  are constants. Define  $X = \sigma Z + \mu$ . Then:

$$f_X = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-(x-\mu)^2/2\sigma^2}$$

That is,  $\sigma Z + \mu$  has the distribution type  $\mathcal{N}(\mu, \sigma^2)$ .

### Derivation of PDF of $\sigma Z + \mu$

Suppose that  $X = \sigma Z + \mu$ . Then:

$$\begin{aligned} F_X(x) &= P[X \leq x] \\ &= P[\sigma Z + \mu \leq x] \\ &= P[Z \leq \frac{x-\mu}{\sigma}] \\ &= \Phi\left(\frac{x-\mu}{\sigma}\right) \end{aligned}$$

Differentiate to find  $f_X$ :

$$\begin{aligned}
f_X(x) &= \frac{d}{dx} F_X(x) \\
&= \frac{d}{dx} \Phi\left(\frac{x-\mu}{\sigma}\right) \\
&= \frac{1}{\sigma} \varphi\left(\frac{x-\mu}{\sigma}\right) \\
&= \frac{1}{\sqrt{2\pi\sigma^2}} e^{-(x-\mu)^2/2\sigma^2}
\end{aligned}$$

From this fact we can infer that  $E[X] = \mu$  and  $\text{Var}[X] = \sigma^2$  whenever  $X \sim \mathcal{N}(\mu, \sigma^2)$ .

### 03 Illustration

#### ☰ Example - Basic generalized normal calculation

Suppose  $X \sim \mathcal{N}(-3, 4)$ . Find  $P[X \geq -1.7]$ .

##### Solution

First write  $X$  as a linear transformation of  $Z$ :

$$X \sim 2Z - 3$$

Then:

$$X \geq -1.7 \iff Z \geq 0.65$$

Look in a table to find that  $\Phi(0.65) \approx 0.74$  and therefore:

$$\begin{aligned}
P[Z \geq 0.65] &\ggg 1 - P[Z \leq 0.65] \\
&\ggg \approx 1 - 0.74 \ggg 0.26
\end{aligned}$$

#### ☰ Example - Gaussian: interval of 2/3

Find the number  $a$  such that  $P[-a \leq Z \leq +a] = 2/3$ .

##### Solution

First convert the question:

$$\begin{aligned}
P[-a \leq Z \leq +a] &\ggg F_Z(a) - F_Z(-a) \\
&\ggg \Phi(a) - \Phi(-a) \\
&\ggg 2\Phi(a) - 1
\end{aligned}$$

Solve for  $a$  so that this value is  $2/3$ :

$$2\Phi(a) - 1 = 2/3 \ggg \Phi(a) = 5/6 \ggg a = \Phi^{-1}(5/6)$$

Use a  $\Phi$  table to conclude  $a \approx 0.97$ .

### ☰ Example - Heights of American males

Suppose that the height of an American male in inches follows the normal distribution  $\mathcal{N}(71, 6.25)$ .

- (a) What percent of American males are over 6 feet, 2 inches tall?
- (b) What percent of those over 6 feet tall are also over 6 feet, 5 inches tall?

#### Solution

(a)

Let  $H$  be a random variable measuring the height of American males in inches, so  $H \sim \mathcal{N}(71, 6.25)$ . Thus  $H \sim 2.5Z + 71$ , and:

$$\begin{aligned} P[H > 74] &\ggg 1 - P[H \leq 74] \\ &\ggg 1 - P[2.5Z + 71 \leq 74] \\ &\ggg 1 - P[Z \leq 1.20] \\ &\ggg 1 - 0.8849 \approx 11.5\% \end{aligned}$$

(b)

We seek  $P[H > 77 \mid H > 72]$  as the answer. Compute as follows:

$$\begin{aligned} P[H > 77 \mid H > 72] &= \frac{P[H > 77]}{P[H > 72]} \\ &\ggg \frac{P[2.5Z + 71 > 77]}{P[2.5Z + 71 > 72]} \\ &\ggg \frac{1 - P[Z \leq 2.4]}{1 - P[Z \leq 0.4]} = \frac{1 - 0.9918}{1 - 0.6554} \approx 2.38\% \end{aligned}$$

### ☰ Example - Variance of normal from CDF table

Suppose  $X \sim \mathcal{N}(5, \sigma^2)$ , and suppose you know  $P[X > 9] = 0.2$ .

Find the approximate value of  $\sigma$  using a  $\Phi$  table.

#### Solution

$$X \sim \mathcal{N}(5, \sigma^2) \implies X \sim \sigma Z + 5$$

So  $1 - P[X \leq 9] = 0.2$  and thus  $P[\sigma Z + 5 \leq 9] = 0.8$ . Then:

$$P[\sigma Z + 5 \leq 9] = P[Z \leq 4/\sigma]$$

so  $P[Z \leq 4/\sigma] = 0.8$ .

Looking in the chart of  $\Phi$  for the nearest inverse of 0.8, we obtain  $4/\sigma = 0.842$ , hence  $\sigma = 4.75$ .